A Lagrangian Half-Quadratic Approach to Robust Estimation and its Applications to Road Scene Analysis

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Abstract

We consider the problem of fitting linearly parameterized models, that arises in many computer vision problems such as road scene analysis. Data extracted from images usually contain non-Gaussian noise and outliers, which makes non-robust estimation methods ineffective. In this paper, we propose an overview of a Lagrangian formulation of the Half-Quadratic approach by, first, revisiting the derivation of the well-known Iterative Re-weighted Least Squares (IRLS) robust estimation algorithm. Then, it is shown that this formulation helps derive the so-called Modified Residuals Least Squares (MRLS) algorithm. In this framework, moreover, standard theoretical results from constrained optimization can be invoked to derive convergence proofs easier. The interest of using the Lagrangian framework is also illustrated by the extension to the problem of the robust estimation of sets of linearly parameterized curves, and to the problem of robust fitting of linearly parameterized regions. To demonstrate the relevance of the proposed algorithms, applications to lane markings tracking, road sign detection and recognition, road shape fitting and road surface 3D reconstruction are presented. Key words: Non Convex problem, Constrained Optimization, Robust

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Estimators, Image Analysis, Stereovision, Linear generative model.

1 1. Introduction

As in many scientific activities, a very common approach to image analysis involves collecting n observations $(x_1, y_1), \dots, (x_n, y_n)$ that take their values in $\mathbb{R}^p \times \mathbb{R}$, and then finding the model that best fits these data. The simplest regression model is the linear one:

$$y_i = X(x_i)^t A + b_i \quad i = 1,n$$
 (1)

where $A = (a_l)_{0 \le l \le d}$ is the vector of (unknown) model parameters, $X(x_i) = (f_l(x_i))_{0 \le l \le d}$ collects the values of some basis of real functions at locations x_i and b_i is the random measurement noise. $X(x_i)$ is also called *design* of the measurement (or experiment) [1]. We assume that the *residuals*, b_i , are independent and identically distributed (i.i.d.), and centered. In real-world applications, residuals are most of the time non-Gaussian and thus some gross errors, called *outliers* may be observed.

Outliers pose a threat on parametric regression under the linear gen-13 erative model (1) in the sense of ordinary Gaussian Maximum-Likelihood 14 estimation. To alleviate their biasing effect on estimation, Huber proposed 15 in [2] to use non-Gaussian Maximum-Likelihood type estimators, a.k.a. M-16 estimators. Bounding the resulting non-quadratic energy with parabolas 17 leads to the so-called Iterative Re-weighted Least Squares (IRLS or IRWLS) 18 or to Modified Residuals Least Squares (MRLS), depending on the analytic 19 form of the parabola. Unfortunately, although the theory of robust estima-20 tion was already known in statistics as early as the mid 80s, it did not cross 21 the barrier of communities. 22

In image reconstruction, the work of the Geman brothers [3–5] for edge-23 preserving regularization was integrated and extended in [6, 7] under the 24 name of Half-Quadratic regularization. The direct connection between M-25 estimation and Half-Quadratic theory was pointed out in [8]. More recently, 26 Lange [9] proposed an interpretation of these algorithms as optimization 27 transfer or bounded optimization algorithms, a particular class of Convex-28 Concave procedures [10] to which the Expectation Maximization (EM) algo-29 rithm also belongs. An interpretation of Half-Quadratic algorithms as EM 30 algorithms was also proposed in [11]. 31

In this paper, we revisit a derivation of Half-Quadratic algorithms which 32 is based on a Lagrangian formalism. The Lagrangian Half-Quadratic ap-33 proach was developed in a series of papers which are here summarized for a 34 unified presentation. Our goal in this article is to show that the Lagrangian 35 Half-quadratic approach offers two non-negligible advantages over existing 36 derivations. First, it eases convergence proofs since we can invoke standard 37 results from the theory of constrained optimization. This is illustrated by 38 the derivation of the classical IRLS algorithm, as presented first in [16] and 39 by the derivation of the less known MRLS algorithm, as first explain here. 40 Second, it helps derive new algorithms, for e.g. simultaneous regression of 41 multiple curves (first presented in [23] and detailed in [12, 13]) or for lin-42 early parameterized region fitting (first presented in [33] for road surface 43 segmentation). Moreover, we show that the MRLS and IRLS algorithms can 44 be used with advantages in many computer vision problems as illustrated 45 with two non-obvious and original application examples in the context of 46 road scene analysis: detection and recognition of traffic signs (previously 47

⁴⁸ presented in [25]) and 3D stereo reconstruction of surfaces (first presented
⁴⁹ in [30] for the road surface case).

The organization of the paper is as follows: in Section 2, the Lagrangian formulation is presented and is applied to the derivation of the IRLS and MRLS algorithms. We also present a family of potential functions, and a continuation heuristic which helps convergence to a satisfactory local minimum. Various applications, involving extensions of the proposed framework, in the field of road scene analysis are proposed in Section 3. We also address difficult problems, such as:

- multiple marking lanes detection and tracking under adverse meteoro logical conditions,
- road sign detection and recognition,
- road pavement detection in images,
- 3D road shape reconstruction using stereovision.

62 2. Robust Parameter Estimation

As explained in [12, 13], the derivation of the robust estimator can be obtained using Lagrange's formulation which leads to the same algorithms as those obtained by the Half-Quadratic and M-estimation approaches. As in these approaches, we focus on a symmetric probability density functions (pdf) of the residual, *b* in the form:

$$p(b) \propto \frac{1}{s} e^{-\frac{1}{2}\phi((\frac{b}{s})^2)} \tag{2}$$

where \propto denotes the equality up to a factor, and *s* is the scale of the noise. With the Maximum-Likelihood Estimation (MLE) method, the problem of estimating the linear model (1) under noise (2) is set as the minimization w.r.t. *A* of the error:

$$e_R(A) = \frac{1}{2} \sum_{i=1}^{i=n} \phi((\frac{y_i - X_i^t A}{s})^2).$$
(3)

To solve this problem, as in the Half-Quadratic approach [4, 7], $\phi(t)$ r3 should fulfill the following hypotheses:

- H0: ϕ is defined and continuous on $[0, +\infty)$ as well as its first and r5 second derivatives,
- **H1**: $\phi'(t) > 0$ (thus ϕ is increasing),

• H2:
$$\phi''(t) < 0$$
 (thus ϕ is concave).

77

⁷⁸ Specifically for the derivation of the MRLS algorithm, a fourth hypothesis ⁷⁹ on $\phi(t)$ is required:

• **H3**:
$$\phi'(t) \leq 1 \ (\phi' \text{ is bounded})$$

As stated in [2], the role of ϕ is to saturate the error in case of a large scaled residual $|b_i| = |y_i - X_i^t A|$, and thus to lower the importance of outliers. The scale parameter, s, sets the residual value from which noisy data points have a good chance of being considered as outliers.

The Lagrangian formulation consists in first rewriting the problem as a constrained optimization problem that is solved by the search for a saddle point of the associated Lagrange function. Then, the algorithm is obtained by alternated minimizations of the dual function. As we will see, for (3), two algorithms can be derived: the first derived algorithm corresponds to the
well-known algorithm named Iterative Re-weighted Least Squares (IRLS) see
Section 2.1, and the second known as the somewhat less popular Modified
Residuals Least Squares (MRLS) see Section 2.2.

93 2.1. Iterative Re-weighted Least Square (IRLS)

First, we rewrite the minimization of $e_R(A)$ as the maximization of $-e_R$. 94 This will allow us later to write $-e_R(A)$ as the extremum of a convex function 95 rather than a concave one, since the negative of a concave function is convex. 96 Second, we introduce the auxiliary variables $w_i = \left(\frac{y_i - X_i^t A}{s}\right)^2$. These variables 97 are needed to rewrite $-e_R(A)$ as the value achieved at the minimum of a 98 constrained problem. This apparent complication is in fact valuable since it 99 allows us to introduce a Lagrange function. Indeed using H1, $-e_R(A)$ can 100 be seen as the value achieved by the minimization w.r.t. $W = (w_i)_{1 \le i \le n}$ of: 101

$$E_{IRLS}(A, W) = \frac{1}{2} \sum_{i=1}^{i=n} -\phi(w_i),$$

¹⁰² subject to *n* constraints $w_i - (\frac{y_i - X_i^t A}{s})^2 \leq 0$, for any *A*. This sub-problem is ¹⁰³ well-posed because it is a minimization of a convex function subject to linear ¹⁰⁴ constraints. Therefore using Kuhn and Tucker's classical theorem [14], if a ¹⁰⁵ solution exists, the minimization of $E_{IRLS}(A, W)$ w.r.t. *W* is equivalent to ¹⁰⁶ search from the unique saddle point of the Lagrangian:

$$L_{IRLS}(A, W, \lambda_i) = \frac{1}{2} \sum_{i=1}^{i=n} -\phi(w_i) + \lambda_i (w_i - (\frac{y_i - X_i^t A}{s})^2),$$

where λ_i are Kuhn-Tucker multipliers ($\lambda_i \geq 0$).

The derivatives of the Lagrange function $L_{IRLS}(A, W, \lambda_i)$ w.r.t. the auxiliary variables W, the unknown variables A, and the Khun-Tucker multipliers λ_i are set to zero to obtain the IRLS algorithm:

111 1. Initialize
$$A^0$$
, and set $k = 1$,

- 112 2. For all $1 \le i \le n$, compute the auxiliary variables $w_i^k = (\frac{y_i X_i^t A^{k-1}}{s})^2$ 113 and the weights $\lambda_i^k = \phi'(w_i^k)$,
- 3. Solve the linear system $\sum_{i=1}^{i=n} \lambda_i^k X_i X_i^t A^k = \sum_{i=1}^{i=n} \lambda_i^k X_i y_i$
- 4. If $||A^k A^{k-1}|| > \epsilon$ then increment k and go to 2, else $A_{IRLS} = A^k$.

As detailed in [13], this algorithm can be derived rigorously from the alternated minimization of the dual error associated to the Lagrange function $L_{IRLS}(A, W, \lambda_i)$. Moreover, it can be shown that such an algorithm always strictly decreases the dual function if the current point is not a stationary point (*i.e.* a point where the first derivatives are all zero) of the dual function [15].

122 2.2. Modified Residuals Least Squares

The IRLS algorithm is into a multiplicative form, but there also exists an 123 algorithm based on an additive Half-Quadratic development called Modified 124 Residuals Least Squares (MRLS), which can be convenient for particular ap-125 plications as illustrated in Section 3.2. The Lagrangian formulation can also 126 be used to derive the MRLS algorithm and to prove its convergence towards a 127 local minimum. Again, we introduce auxiliary variables and rewrite $e_R(A)$ as 128 the value achieve as the search for a saddle point of the associated Lagrange's 129 function. Then, the algorithm is obtained by alternated minimization of the 130 dual function. 13

Let us introduce function $f(b) = b^2 - \phi(b^2)$ defined on $[0, +\infty[$, and thus rewrite $e_R(A)$ as:

$$e_R(A) = \frac{1}{2} \sum_{i=1}^{i=n} \left(\frac{y_i - X_i^t A}{s}\right)^2 - f\left(\left|\frac{y_i - X_i^t A}{s}\right|\right).$$

This will allow us to later write $e_R(A)$ as the value achieved at the mini-134 mum of a convex function for any A. First, we have $f'(b) = 2b(1 - \phi'(b^2))$ 135 and thus f is increasing as required by H3. Moreover, we have f''(b) =136 $2(1-\phi'(b^2)-2\phi''(b^2)b^2)$, and thus $f''(b) \ge 0$ using **H2** and **H3**. Second, we 137 introduce the auxiliary variables $\omega_i = \left|\frac{y_i - X_i^t A}{s}\right|$. These variables are needed to 138 rewrite $e_R(A)$ as the value achieved at the minimum of a constrained problem. 139 Indeed, using the fact that f is increasing, the second term $\sum_{i=1}^{i=n} f(|\frac{y_i - X_i^t A}{s}|)$ 140 of $e_R(A)$ can be seen as the value achieved at the minimization with re-141 spect to $\Omega = (\omega_i)_{1 \le i \le n}$ of $E_{MRLS}(A, \Omega) = \sum_{i=1}^{i=n} f(\omega_i)$, subject to n con-142 straints $\left|\frac{y_i - X_i^t A}{s}\right| - \omega_i \leq 0$. This last sub-problem is well-posed because it is 143 a minimization of a convex function subject to linear constraints w.r.t. Ω . 144 Therefore using the classical Kuhn and Tucker's theorem [14], if a solution 145 exists, the minimization of $E_{MRLS}(A, \Omega)$ with respect to Ω is equivalent to 146 the search of the unique saddle point of the Lagrangian of the sub-problem: 147

$$L_{MRLS}(A, \Omega, \lambda_i) = \sum_{i=1}^{i=n} f(\omega_i) + \lambda_i \left(\left| \frac{y_i - X_i^t A}{s} \right| - \omega_i \right)$$

¹⁴⁸ More formally, we have proved for any A:

$$\sum_{i=1}^{i=n} f(|\frac{y_i - X_i^t A}{s}|) = \min_{\omega_i} \max_{\lambda_i} L_{MRLS}(A, \Omega, \lambda_i).$$
(4)

¹⁴⁹ Using the saddle point property, we can change the order of variables ω_i ¹⁵⁰ and λ_i in (4). $L_{MRLS}(A, \Omega, \lambda_i)$ being convex with respect to Ω , it is equivalent to search for a minimum of $L_{MRLS}(A, \Omega, \lambda_i)$ with respect to Ω and to have its first derivatives zero. Thus, we deduce $\lambda_i = f'(\omega_i)$. Recalling that f' is increasing, this last equation can be used to substitute ω_i in L:

$$\sum_{i=1}^{n} f(|\frac{y_i - X_i^t A}{s}|) = \max_{\lambda_i} L_{MRLS}(A, f'^{-1}(\lambda_i), \lambda_i)$$
(5)

Therefore, we deduce that the original problem is equivalent to the followingminimization:

$$\min_{A} e_R(A) = \frac{1}{2} \min_{A,\lambda_i} \sum_{i=1}^{i=n} (\frac{y_i - X_i^t A}{s})^2 - L_{MRLS}(A, f'^{-1}(\lambda_i), \lambda_i)$$

The function $\mathcal{E}(A, \lambda_i) = \frac{1}{2} \sum_{i=1}^{i=n} (\frac{y_i - X_i^i A}{s})^2 - \frac{1}{2} L_{MRLS}(A, f'^{-1}(\lambda_i), \lambda_i)$ is the dual error of the original problem. Notice that the dual error \mathcal{E} is rather simple with respect to A, contrary to the original error e_R . Indeed, the dual error can be rewritten as:

$$\mathcal{E}(A,\lambda_i) = \frac{1}{2} \sum_{i=1}^{i=n} (|\frac{y_i - X_i^t A}{s}| - \frac{\lambda_i}{2})^2 + \xi(\lambda_i)$$

with $\xi(\lambda_i) = -\frac{\lambda_i^2}{4} - f(f'^{-1}(\lambda_i)) + \lambda_i f'^{-1}(\lambda_i)$

Taking its second derivatives, we deduce that \mathcal{E} is convex with respect to *A.* \mathcal{E} is also convex with respect to λ_i since f is convex and $\frac{\partial^2 \mathcal{E}}{\partial \lambda_i^2} = \frac{1}{f''(f'^{-1}(\lambda_i))} \geq$ 0. However, when $e_R(b^2)$ is not convex, \mathcal{E} may not be convex with respect to both A and λ_i . In such a case, $\mathcal{E}(A, \lambda_i)$ does not have a unique minimum.

An alternate minimization of the dual error \mathcal{E} , with respect to A and each λ_i , involves the sign of $y_i - X_i^t A$, and leads to the modified residual iterative algorithm, already derived in the Half-Quadratic and M-estimators approaches: 169 1. Initialize A^0 , and set k = 1,

2. For all
$$1 \le i \le n$$
, compute the auxiliary variables $\omega_i^{sign,k} = \frac{y_i - X_i^t A^{k-1}}{s}$
and the weights $\lambda_i^{sign,k} = \omega_i^{sign,k} (1 - \phi'(\omega_i^{sign,k^2})),$

3. Solve the linear system $\sum_{i=1}^{i=n} X_i X_i^t A^k = \sum_{i=1}^{i=n} X_i (y_i - s \lambda_i^{sign,k}),$

4. If
$$||A^k - A^{k-1}|| > \epsilon$$
 then increment k and go to 2, else $A_{MRLS} = A^k$.

In this algorithm we introduce the two notations $\omega_i^{sign,k} = \omega_i^k sign(y_i - \omega_i^k)$ 174 $X_i^t A^{k-1}$) and $\lambda_i^{sign,k} = \lambda_i^k sign(y_i - X_i^t A^{k-1})$, which are signed version of the 175 auxiliary variables ω_i and λ_i . Like in the previous section, it can be shown 176 that this alternate minimization of \mathcal{E} always strictly decreases the dual func-177 tion, if the current point is not already a stationary point [15]. Following the 178 lines of [12], it can then be deduced that the proposed algorithm is globally 179 convergent, i.e., it converges towards a local minimum of $e_R(A)$ for all initial 180 A_0 's which are not a maximum or a saddle point of $e_R(A)$. 181

182 2.3. Non-Gaussian Noise Model

We are interested in a parametric family of functions for noise modeling, in the form of (2) that allows a continuous transition between different kinds of useful probability distributions. We thus focus on a simple parametric probability density functions (pdf) of the form:

$$S_{\alpha,s}(b) \propto \frac{1}{s} e^{-\frac{1}{2}\phi_{S_{\alpha}}((\frac{b}{s})^2)} \tag{6}$$

where the associated ϕ function is $\phi_{S_{\alpha}}(t) = \frac{1}{\alpha}((1+t)^{\alpha}-1)$. This is the so-called Smooth Exponential Family (SEF) introduced in [12, 16, 17] which is suitable for the IRLS and MRLS algorithms, since (for $\alpha < 1$) it satisfies the four hypotheses **H0**, **H1**, **H2** and **H3**.



Figure 1: SEF noise models, $S_{\alpha,s}$ for s = 1 and different values of α . Notice how tails are getting heavier as α decreases.

The parameter α allows a continuous transition between well-known statistical laws such as Gauss ($\alpha = 1$), smooth Laplace ($\alpha = \frac{1}{2}$) and Geman & McClure [18] ($\alpha = -1$). These laws are shown on Figure 1. Note that, for $\alpha < 0$, $S_{\alpha,s}$ can be always normalized on a bounded support, so it can still be considered as a pdf. In the smooth exponential family, when α is decreasing, the probability to observe large, not to say very large errors (outliers), increases.

In the IRLS and MRLS algorithms, the weights λ_i are a function of $\phi'(t)$.

For the SEF, this function is simply $\phi'_{S_{\alpha}}(t) = (1+t)^{\alpha-1}$. Notice that while the pdf is not defined when $\alpha = 0$, the corresponding $\phi'(t)$ does and that it corresponds in fact to the Cauchy distribution or Student's t-distribution [17]. It can be shown, see [12, 17], that the breakdown point of SEF estimators

is increasing towards the maximum achievable value, that is 50%, as $\alpha \in [0, 0.5]$ decreases. The maximum goes to 50% when $\alpha \to 0$.

205 2.4. Graduated Non-Convexity (GNC)

The function $\phi'(t)$, used in the IRLS and MRLS algorithms, becomes 206 more sharply peaked and heavily tailed as α decreases. As a consequence, 207 the lower the α , the lower the effect of outliers on the result. Therefore, 208 the algorithm produces a more robust fitting. However, when α decreases, 209 the error function $e_R(A)$ becomes less and less smooth. If $\alpha = 1$, the cost 210 function is a paraboloid and thus there exists a unique global minimum. For 211 α between 1 and 0.5, the cost function is convex w.r.t. A and thus there still 212 exists a unique global minimum. By decreasing α to values lower than $\frac{1}{2}$, 213 local minima appear. This is illustrated in Figure 2 where the error function 214 $e_R(A)$ is shown for four decreasing values of α . 215

Following the principle of the GNC method [19], the localization property 216 of the robust fit w.r.t. the decreasing parameter α can be used to converge 217 toward a local minimum close to the global one. Convexity is first enforced 218 using $\alpha = 1$ or $\alpha = 0.5$. Then, a sequence of fits with a stepwise decreasing 219 α is performed in continuation, *i.e.* each time using the current output fit 220 as an initial value for the next fitting step. Of course, as detailed in [20], 221 α must be decreased slowly, otherwise the curve fitting algorithm might be 222 trapped into a shallow local minimum far from the global one. 223



Figure 2: The error function $e_R(A)$ for an example of scalar data with two clusters, for different values of α . Notice the progressive appearance of the second minimum while α decreases.

224 3. Applications and extensions in road scene analysis

225 3.1. Lane-markings Fitting

In our approach of lane marking detection, road marking elements are first extracted from a road scene image using one of the algorithm compared in [21]. Then, the centers of marking elements are fitted with a curve using an IRLS algorithm. As explained in [22], it is possible to use different curve families to model the shape of a lane marking, such as polynomial curves or hyperbolic polynomial curves, which better fit road edges on long rangedistances.



Figure 3: Fitting on a real image assuming (from left to right, up to down) Gauss, smooth Laplace, Cauchy, and Geman & McClure [18] noise pdfs. Data points are shown in blue, yellow lines are the initial A^0 and green lines are the fitting results.

Figure 3 illustrates the importance of robust fitting in images with many outliers. The yellow line depicts the initial A^0 obtained from the previous image. The green ones are the fitting results A_{IRLS} , assuming respectively Gauss, smooth Laplace, Cauchy, and Geman & McClure [18] noises (as defined in Section 2.3). A correct fit is achieved only with the last two pdf's, which correspond to non-convex errors $e_R(A)$. The scale *s* is fixed to 4 pixels



²³⁹ from the analysis of the residuals on a ground-truth database.

Figure 4: Fitting of two lines along a sequence with rain and visibility difficulties.

As visible on Figure 3, several lane markings can be seen in a road image. We formulate the problem of robustly estimating in a simultaneous fashion m linearly parameterized curves A_j , whose parameters may be concatenated in a single vector $A = (A_j)_{j=1...m}$, as the optimization of:

$$e_{MR}(A) = -\sum_{i=1}^{i=n} \log \sum_{j=1}^{j=m} e^{-\frac{1}{2}\phi((\frac{y_i - X_i^t A_j}{s})^2)}.$$
(7)

The Lagrangian formalism eases deriving from the minimization of (7) the so-called Simultaneous Robust Multiple Fitting (SRMF) algorithm [12]:

²⁴⁶ 1. Choose an initial estimate A^0 , and initialize loop index to k = 1.

247 2. For each
$$1 \leq j \leq m$$
, compute vector $W_j^k = \frac{1}{s}(X^t A_j^{k-1} - Y)$
248 and matrix $B_j^k = diag \left\{ \frac{e^{-\phi\left(w_{ji}^k\right)}}{\sum_{j=1}^{j=m} e^{-\phi\left(w_{ji}^k\right)}} \phi'\left(w_{ji}^k\right) \right\}_{i=1\dots n}$.
249 3. For each $1 \leq j \leq m$, solve $XB_i^k X^t A_j^k = XB_i^k Y$.

4. If
$$||A^k - A^{k-1}|| > \epsilon$$
 then increment k and go to 2, else $A_{SRMF} = A^k$.

With notations $X = (X_i)_{1 \le i \le n}$, $Y = (y_i)_{1 \le i \le n}$ and $W_j^k = (w_{ji}^k)_{1 \le i \le n}$. More details on the derivation and on the use of the SRMF algorithm may be found in [12], where the convergence to a local minimum is proved. Notice that the IRLS algorithm may be seen as a special case of the SRMF algorithm with m = 1. Moreover, note that the SRMF weights incorporate a probability ratio, which is customary used in classification algorithms as a membership function.

The tracking of several curves along a sequence of images is performed by embedding the SRMF algorithm within a Kalman filter (see [13, 16, 23] for details) as illustrated in Figure 4.

²⁶¹ 3.2. Appearance-based road sign detection and recognition

Appearance-based models have met unquestionable success in the field of 262 object detection and recognition, since the early 1990s [24]. In such mod-263 els, objects are represented by their raw brightness or color vector, without 264 any feature extraction. To capture non-linear appearance variations while 265 using a linear generative model, we proposed in [25, 26] an approach based 266 on probabilistic principal component analysis [27]. More specifically, any ob-267 served image y may be reconstructed on a relatively low-dimensional basis of 268 functions, that span the eigenspace of the covariance matrix obtained from 269 a set of representative images: $y = \mu + X^t A + b$, where μ is the sample mean 270

of the learning set. This is a special case of model (1) where the X_i 's are 271 not analytical but *statistical* basis functions. Figure 5 shows an example of 272 sample images and the corresponding basis vectors. Detection and recogni-273 tion tasks involve estimating the *latent variable*, A. Such a regression task 274 may be performed robustly in the Half-Quadratic framework. In this par-275 ticular case, the scale variable, s, is fixed accordingly to a statistical study 276 of the residuals obtained by reconstructing the (outlier-free) learning image 27 samples [25]. As shown in Figure 5, robust estimation provides both a cor-278 rect reconstructed image and outliers map. Note that, in this particular case 279 where X is orthogonal and the dimension of the eigenspace is typically about 280 20, the MRLS algorithm is recommended since it can be more computation-281 ally efficient than the IRLS. We shown in [25] that enforcing a well-suited 282 prior to the latent variable is beneficial to the quality of detection. Moreover, 283 we proposed in [26] an original algorithm which handles both robust likeli-284 hoods and arbitrary non-parametric priors on A in the framework of Mean 285 Shift [28] (which, besides, is also a Half-Quadratic algorithm [29]). This 286 approach shows very good performance in recognition task on real images. 28

288 3.3. Road profile 3D reconstruction

Another application of our Lagrangian formalism is road profile reconstruction from stereovision, assuming a rectified geometry. In practice, the left and right views mostly contain the image of the road. The surface of the road in front of the vehicle defines a mapping between the two images. Assuming a polynomial model for the road profile [30], within the disparity space [31], this mapping between a point (u, v) in the left image and a point



Figure 5: First row: 7 of the N = 43 images used for learning. Second row: mean image and first 3 eigenvectors obtained by principal component analysis from the learning database. Third row: recognition experiment. From left to right: analyzed image, reconstruction from robustly estimated co-ordinates in eigenspace, outliers map, recognized image.

²⁹⁵ (u', v') in the right image is:

where $X(v) = (1, v, v^2, \dots, v^n)^t$ is the vector of monomials and $A = (a_0, \dots, a_n)$ is the vector of unknown parameters related to the vertical road profile.

The 3D reconstruction of the road surface vertical profile needs horizontal correspondences between the edges on the road in the left and right images. Instead of performing first the matching and then the 3D reconstruction, we solve the problem by formulating it as the fitting of linear model A on the set of all possible horizontal correspondences ((u, u'), v). The robustness ³⁰³ of the IRLS estimator is thus of major importance, since many of these
 ³⁰⁴ correspondences are wrong.

More specifically, we formulate in [30] the estimation problem using a Generalized M-estimator as:

$$e_R(A) = \frac{1}{2} \sum_{(u,u'),v} d((u,u'),v)\phi((\frac{u+X(v)^t A - u'}{s})^2),$$
(9)

where d((u, u'), v) measures the local similarity in terms of gray or color gradients between the two matched edge pixels (u, v) and (u', v). The IRLS algorithm resulting from our Lagrangian derivation features an extra factor in the weights, which becomes in that case $\lambda_{(u,u'),v} = d((u, u'), v)\phi'((\frac{u+X(v)^t A - u'}{s})^2)$

To improve convergence towards a local minima close to the global one, the optimization is iterated for several scales s. At the beginning, s is set to a large value and is then decreased step by step to 1 pixel. This is another kind of GNC based on the use of the scale s rather than on the use of α as described in Section 2.4.

It is important to notice that, in the obtained algorithm, the matching 316 is one-to-several rather than one-to-one. As experimented in [32], one-to-317 several correspondence provides better convergence towards an interesting 318 local minimum, and thus outperforms Iterative Closest Point or other one-319 to-one correspondence algorithms. This improved convergence property is 320 detrimental in terms of computational burden. To avoid penalizing the al-321 gorithm, we perform a matching decimation in the spirit of [32]: correspon-322 dences such as u + X(v)A - u' is larger than 3s are discarded. This decimation 323 is performed only one time at each scale s without a significant lost of accu-324 racy, see [30]. Figure 6 displays the edges of the right image obtained after 325



Figure 6: Edges of the right image obtained after alignment on the left image using road profile fitting.

alignment on the left image, showing the quality of the 3D reconstruction of
the road vertical profile on four consecutive images.

328 3.4. Road region fitting

Thanks to the Lagrange's formulation, we proposed in [33] a new algorithm for region fitting which is robust to missing data. It assumes linearly parameterized borders and it is applied to the modeling of the road region in a road scene image. As in Section 3.1, we use polynomial or hyperbolic polynomial curves to model the borders of the road region. In vector notations,



Figure 7: From left to right, first row: original image, color segmentation of the road, resulting region fit on the segmentation, second row: left part of the region, right part of the region, complete resulting region.

right and left borders are thus written as $u_r = X(v)^t A_r$ and $u_l = X(v)^t A_l$. This region fitting algorithm is obtained as the minimization of the following classical least squares error:

$$e_{RRF}(A_l, A_r) = \int \int_{Image} \left[P(u, v) - \mathcal{R}_{A_l, A_r}(u, v) \right]^2 du dv \tag{10}$$

between the image P(u, v) which corresponds to the probability of each pixel 337 to belong to the "road" class and the function $\mathcal{R}_{A_l,A_r}(u,v)$ between 0 and 1 338 modeling the road region. The probability map is obtained by segmentation 339 of the road colors as detailed in [34], see second image, first row in Figure 7 for 340 an example. This road region is parameterized by A_l and A_r , the parameters 341 of the left and right border curves. Noticing that function $u - X(v)^t A$ is 342 defined for all pixel coordinates (u, v), this function is negative on the left 343 of the curve and positive on its right. We thus can use it to define region 344

345 \mathcal{R}_{A_l,A_r} as:

$$\mathcal{R}_{A_l,A_r}(u,v) = \left(g\left(\frac{X(v)^t A_l - u}{s}\right) + \frac{1}{2}\right) \left(\frac{1}{2} - g\left(\frac{X(v)^t A_r - u}{s}\right)\right)$$
(11)

where g is an increasing odd function with $g(+\infty) = \frac{1}{2}$. The s parameter tunes the strength of smoothing on the road region borders. The second row in Figure 7 display an example of the first factor in (11), of the second term, and of the result \mathcal{R}_{A_l,A_r} .

By substitution of the previous model in (10), we rewrite it in its discrete form as:

$$e_{RRF}(A_l, A_r) = \sum_{ij \in Image} \left[P_{ij} - \left(g\left(\frac{X_i^t A_l - j}{s}\right) + \frac{1}{2} \right) \left(\frac{1}{2} - g\left(\frac{X_i^t A_r - j}{s}\right) \right) \right]^2 (12)$$

Again, we apply the Lagrange's formulation, which allows us to derive the associated iterative algorithm. As detailed in [33], to introduce constrains, the even function g^2 of the left and right residuals is rewritten as $g^2(t) = h(t^2)$, after expansion of the square in (12). Thus auxiliary variables $\omega_{ij}^l = \left(\frac{X_i^t A_l - j}{s}\right)$, $\omega_{ij}^r = \left(\frac{X_i^t A_r - j}{s}\right), \ \nu_{ij}^l = \left(\frac{X_i^t A_l - j}{s}\right)^2$ and $\nu_{ij}^r = \left(\frac{X_i^t A_r - j}{s}\right)^2$ are introduced. The associated Lagrange function is:

$$L_{RRF} = \sum_{ij} \left[h(\nu_{ij}^{l})h(\nu_{ij}^{r}) + \frac{1}{4}(h(\nu_{ij}^{l}) + h(\nu_{ij}^{r})) + (2P_{ij} - 1)g(\omega_{ij}^{l})g(\omega_{ij}^{r}) \right. \\ \left. + (P_{ij} - 1/4) \left[-g(\omega_{ij}^{l}) + g(\omega_{ij}^{r}) \right] - h(\nu_{ij}^{l})g(\omega_{ij}^{r}) + h(\nu_{ij}^{r})g(\omega_{ij}^{l}) \right] \\ \left. + \sum_{ij} \lambda_{ij}^{l} \left(\omega_{ij}^{l} - \frac{X_{i}^{t}A_{l} - j}{\sigma} \right) + \lambda_{ij}^{r} \left(\omega_{ij}^{r} - \frac{X_{i}^{t}A_{r} - j}{\sigma} \right) \right. \\ \left. + \sum_{ij} \mu_{ij}^{l} \left(\nu_{ij}^{l} - \left(\frac{X_{i}^{t}A_{l} - j}{\sigma} \right)^{2} \right) + \mu_{ij}^{r} \left(\nu_{ij}^{r} - \left(\frac{j - X_{i}^{t}A_{r} - j}{\sigma} \right)^{2} \right) \right.$$
(13)

The derivatives of (13) w.r.t.: the auxiliary variables, the unknown variables A_l and A_r , and the Khun-Tucker multipliers λ_{ij}^l , λ_{ij}^r , μ_{ij}^l μ_{ij}^r are set to zero to obtain the algorithm. Like in the previous section, to improve convergence towards a local minima close to the global one, the optimization is iterated for several scales s.

The last images in first and second rows of Figure 7 show the resulting fit. We observed that fitting algorithms based on a region model are more robust to outliers compared to algorithm based on edges.

366 4. Conclusion

We presented the Lagrange's formulation of Half-Quadratic approach and 367 showed how it can be used to derive the classical IRLS and MRLS algorithms 368 for robustly estimating the parameters of a linear generative model, and to 369 prove their convergence. Examples taken from road scene analysis illustrated 370 the interest of these algorithms in difficult real-world applications such as 37 lane marking detection and tracking or road sign detection and recognition. 372 Other examples, namely simultaneous multiple lane fitting, road profile re-373 construction and road region detection, demonstrated the flexibility of our 374 framework which helps in the derivation of new algorithms. The obtained 375 algorithms are quite fast: it takes around 0.2 s to fit simultaneously multiple 376 lanes, around 0.5 s to reconstruction the road profile, and around 3 s to fit 377 the road region, using a Core2 Duo processor on images of size 640×512 . 378

One may notice that the Lagrange's formulation of Half-Quadratic approach has interesting connections with the Expectation Maximization (EM) approach which will be interesting to investigate more in the future.

One difficulty remains to be solved, which is of practical importance in the context of road scene analysis since correlation matrices are useful for validating the fits. In all our derivation, we supposed independent noise. Our experience is that this assumption, which seems sufficient as long as only fitting is concerned, is not pertinent any more when estimating the correlation matrix of fits in practical applications. So the question is: how would it be possible to introduce correlated noise with heavy-tailed distribution within the Half-Quadratic approach?

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